# CE 269 Traffic Engineering 

## Lecture 17 <br> Introduction to Queueing Theory

## Previously on Traffic Engineering

We can easily analyze delays at a junction in the case of deterministic arrivals and departures. This is an example of $D / D / 1$ queues. More general settings will be discussed later in the course.

Suppose the arrival rate of vehicles at an approach be $v$ and saturation flow rate be $s$. If the cycle time is $C$,

- The number of vehicles arriving at the junction in one cycle is $v C$.
- The maximum number of vehicles that can leave is $s g$, where $g$ is the effective green.


## Previously on Traffic Engineering

Suppose $s g>v C$. Calculate the following quantities:


- The time to clear the queues after the start of the effective green.

$$
t_{c}=\frac{v r}{s-v}
$$

- The proportion of the cycle time with a queue.

$$
t_{c}=\frac{r+t_{c}}{C}
$$

- Total vehicle delay per cycle and average delay per vehicle.

$$
D_{t}=\frac{v r^{2}}{2(1-v / s)} D_{\text {avg }}=\frac{0.5 C(1-g / C)^{2}}{1-(v / c)(g / C)}
$$

## Previously on Traffic Engineering

The earlier examples makes a few implicit assumptions.

- Arrival rates are uniform, which is not always true. Vehicles at isolated intersection usually follow a random arrival process. If the signal is part of a coordinated network, then arrivals are in batches or platoons.
- The queues are assumed to stack vehicles on top of one another. This is also called a point-queue model.

Also, in practice traffic can switch between under-saturated and oversaturated conditions. To address this issue, the delay is usually broken down into uniform delay and overflow delay.

## Previously on Traffic Engineering


(a) Stable Flow

(b) Individual Cycle Failures Within a Stable Operation

(c) Demand Exceeds Capacity for a

Significant Period

## Lecture Outline

1 Review of Probability Distributions
2 Continuous Time Markov Chains
3 Single Server Queues
4 Multiple Servers and Network of Queues

## Lecture Outline

## Review of Probability Distributions

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## Poisson Distribution

The Possion random variable is used to count the number of random events in a time period.

For example,

- The number of accidents that occur on a highway in an year.
- The number of customers served by a teller at a bank.

It is assumed that occurrences of events are independent of each other. Also, we assume that two or more events cannot happen simultaneously. Further, the average rate of occurrence of events is known and assumed constant.
https://www.randomservices.org/random/apps/PoissonExperiment. html

## Review of Probability Distributions

## Poisson Distribution

## Definition

PMF of a Poisson distributed random variable with parameter $\lambda>0$ is

$$
\mathbb{P}(X=x)=p_{X}(x)=\frac{e^{-\lambda} \lambda^{x}}{x!}
$$

## Claim

Suppose $X \sim \operatorname{Pois}(\lambda), \mathbb{E}(X)=V(X)=\lambda$



## Review of Probability Distributions

## Exponential Distribution

Exponential distribution is commonly used to model time between consecutive events when the events occur according to Poisson distribution.

For example,

- The time duration between two accidents on a highway
- The amount of time taken by a bank teller to serve a customer
- The time between two arrivals at a checkout queue


## Exponential Distribution

## Probability Density Function

## Definition

Suppose $X \sim \exp (\lambda)$, its probability density function is defined as

$$
f_{X}(x)= \begin{cases}\lambda e^{-\lambda x} & \text { if } x \geq 0 \\ 0 & \text { otherwise }\end{cases}
$$

## Claim

Suppose $X \sim \exp (\lambda)$, its CDF is

$$
F_{X}(x)=1-e^{-\lambda x}
$$

Check if $d F_{X}(x) / d x=f_{X}(x)$

## Exponential Distribution

## PMF and CDF

PDF and CDF of an exponentially distributed random variable with $\lambda=2$ are shown below.


## Claim

If $X \sim \exp (\lambda)$, then $\mathbb{E}(X)=1 / \lambda$ and $V(X)=1 / \lambda^{2}$

## Exponential Distribution

## Connections with Poisson Distribution

To see why inter-arrival times of an Poisson distributed random variable is exponentially distributed, let $X \sim \operatorname{Pois}(\lambda)$.

Consider a time window $t$. The probability that there are zero arrivals in $t$ is given by

$$
\mathbb{P}(X=0)=\frac{(\lambda t)^{0} e^{-\lambda t}}{0!}=e^{-\lambda t}
$$

If $Y$ is the inter-arrival time, then $\mathbb{P}(X=0)=\mathbb{P}(Y>t)$.
Hence, $\mathbb{P}(Y \leq t)=1-e^{-\lambda t}$, which is the CDF of the exponential random variable.

## Exponential Distribution

## Memoryless property

Suppose that the inter-arrival times of buses at a bus stop are exponentially distributed with rate $\lambda$. Let $X$ be the arrival time of the next bus.

Assuming, that you have been waiting for $t$ minutes (right after the passing of the previous bus), what is the probability that you will have to wait at least another $s$ minutes.

$$
\mathbb{P}(X>s+t \mid X>t)=?
$$

For this reason, exponential random variable is said to exhibit a memoryless property.

## Lecture Outline

## Continuous Time Markov Chains

## Continuous Time Markov Chains

## Discrete-Time Markov Chains

## Definition (Markov Property)

A stochastic process $\left\{X_{n}, n \geq 0\right\}$ with a countable state space $S$ is called a DTMC if $\forall n \geq 0, i, j \in S$,

$$
\mathbb{P}\left[X_{n+1}=j \mid X_{n}=i, X_{n-1}, X_{n-1}, \ldots, X_{0}\right]=\mathbb{P}\left[X_{n+1}=j \mid X_{n}=i\right]
$$

## Definition (Time Homogeneity)

A DTMC $\left\{X_{n}, n \geq 0\right\}$ is said to be time homogeneous if $\forall n \geq 0, i, j \in S$,

$$
\mathbb{P}\left[X_{n+1}=j \mid X_{n}=i\right]=p_{i j}
$$

i.e., RHS does not depend on $n$ or $p_{i j}(n)=p_{i j} \forall n \geq 0$

## Continuous Time Markov Chains

## Discrete-Time Markov Chains

The probability with which the system moves from $i$ to $j, p_{i j}$, is called the transition probability and the matrix of $p_{i j}$ values is called the one-step transition probability matrix.

$$
P=\left[p_{i j}\right]_{|S| \times|S|}
$$

Note that $P$ can have countably infinite rows and columns.

## Definition (Stochastic Matrix)

A square matrix $P=\left[p_{i j}\right]_{|S| \times|S|}$ is called right stochastic if
$1 p_{i j} \geq 0 \forall i, j \in S$
2 $\sum_{j \in S} p_{i j}=1 \forall i \in S$

Transition matrices of a Markov chain are right stochastic matrices.

## Continuous Time Markov Chains

## Discrete-Time Markov Chains

The transition probability matrix can also be visualized as a directed graph in which the states are nodes and an arc $(i, j)$ exists only if $p_{i j}>0$.
$\left.P=\begin{array}{c} \\ 1 \\ 2 \\ 3\end{array} \begin{array}{ccc}1 & 2 & 3 \\ 0.1 & 0.2 & 0.7 \\ 0.6 & 0 & 0.4 \\ 0.4 & 0 & 0.6\end{array}\right]$

http://setosa.io/ev/markov-chains/
The $P$ matrix alone doesn't fully describe a DTMC. We'd also need to know the initial distribution.

$$
a_{i}=\mathbb{P}\left[X_{0}=i\right] \forall i \in S
$$

Let $a$ be row vector of $a_{i}$ 's. A Markov chain can thus be fully specified using ( $S, P, a$ ).

## Continuous Time Markov Chains

## Definition (CTMC)

Suppose $\{X(t), t \geq 0\}$ is a continuous time stochastic process with a countable state space $S=\{0,1, \ldots\}$. Then $X($.$) is a CTMC if the$ transition probabilities have the following Markov property

$$
\mathbb{P}[X(s+t)=j \mid X(r): 0 \leq r \leq s]=\mathbb{P}[X(s+t)=j \mid X(s)]
$$

We can think of transitions in CTMCs in two steps. Imagine that the system spends an exponential amount of time in each state and when it transitions, it does like a DTMC.

Instead of dealing with a transition matrix, we define a rate matrix whose $(i, j)$ th element indicates the product of the rate at which we leave state $i$ and the probability of transitioning to $j$.

## Continuous Time Markov Chains

## Example

Consider a two-state CTMC with the following transition rates. Imagine that it represents a machine replacement problem.


The rate matrix is written as $\left[\begin{array}{cc}-\lambda & \lambda \\ \mu & -\mu\end{array}\right]$

## Continuous Time Markov Chains

## Example

One metric of interest is the steady state probability $p_{j}$ that the system is in state $j$.

## Definition (Fundamental Theorem of CTMC)

Let $\{X(t), t \geq 0\}$ be an irreducible CTMC. The steady state probabilities are a solution to the following equations

$$
\begin{gathered}
p Q=0 \\
\sum_{j \in S} p_{j}=1
\end{gathered}
$$

These are also called global balance equations as they can be written as the rate out $=$ rate in for every state.

Find the steady state probabilities for the machine reliability example.

## Lecture Outline

## Single Server Queues

## Single Server Queues

## Queuing Theory

Queuing Theory involves studying quantities such as queue lengths and waiting times in a single queue or a network of queues. Application areas include:

- Checkout lines, ATMs
- Call centers
- Traffic signals and Toll plazas
- Airports
- Communication and Internet
- Manufacturing


## Single Server Queues

## Queuing Theory

Consider a single queue with an arrival process $A(t)$ and departure process $D(t)$.


The horizontal and vertical cross-sections tell us the amount of time spent in the queue and the queue length.

The area represents the overall time spent by everyone in the queue.

## Single Server Queues

Let us now analyze a special case where the arrival process is a Poisson process with rate $\lambda$ and the service times are exponential with rate $\mu$. The fraction $\rho=\lambda / \mu$ is called the traffic intensity.


Can you write the global balance equations and find the steady state probabilities?

## Single Server Queues

Poisson Arrivals and Exponential Service Times
Thus, the steady state probabilities are

$$
\begin{gathered}
p_{0}=1-\rho \\
p_{i}=\rho^{i}(1-\rho)
\end{gathered}
$$

How do we find the expected length of the queue?

$$
\mathbb{E}(L)=\sum_{j=0}^{\infty} j p_{j}=\sum_{j=0}^{\infty} j \rho^{j}(1-\rho)=\frac{\rho}{1-\rho}
$$

We can also find the expected waiting time by defining new random variables and taking their expectations.

$$
\mathbb{E}(T)=\frac{1}{\mu-\lambda}
$$

How are $\mathbb{E}(L)$ and $\mathbb{E}(T)$ related? $\mathbb{E}(L)=\lambda \mathbb{E}(T)$. This is called Little's law and holds in more general settings.

## Lecture Outline

## Multiple Servers and Network of Queues

## Multiple Servers and Network of Queues

## Kendall Notation

The example that we discussed so far is referred to as $M / M / 1$ queues. In general, the assumptions used in modelling a queue are represented using Kendall's notation

$$
A / S / c / K
$$

- A Arrival process
- $S$ Service time distribution
- $c$ Number of servers
- K Capacity of the queue
$M$ indicates Markovian or memoryless and is used to indicate Poisson processes and exponential service times.Other options include $D$ for deterministic and $G$ for general distribution.


## Multiple Servers and Network of Queues

## Queue Discipline

Additionally, we may also specify the service discipline.

- FIFO
- LIFO
- Random Service
- Round Robin
- Priority Order


## Multiple Servers and Network of Queues

## M/M/2 queue

Imagine a single queue but two servers with difference service rates.


Systems like these can also be analysed using CTMCs

## Multiple Servers and Network of Queues



The expressions are relatively simpler when the service rates are the same. For an $M / M / c$ queue, let $\rho=\frac{\lambda}{c \mu}$, then $p_{0}$ is

$$
\left[\left(\sum_{k=0}^{c-1} \frac{(c p)^{k}}{k!}+\frac{(c p)^{c}}{c!} \frac{1}{1-p}\right)\right]^{-1}
$$

## Multiple Servers and Network of Queues

## Network of Queues

For a network of queues, we have to resort to simulation. Software packages such as Anylogic, Arena can be used for this purpose.

https://www.supositorio.com/rcalc/rcalclite.htm

## Your Moment of Zen

v horld +21
Queue is just Q followed by 4 silent letters.

- Robot_Spider
$+62$
They aren't silent. They're waiting their turn.

